Cognitive Control in Preadolescent Children With Risk Factors for Metabolic Syndrome

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Objective: To investigate the relationship between cognitive control and metabolic syndrome (MetS) risk factors in preadolescent children while controlling for aerobic fitness and weight status. Methods: Hierarchical regression analyses were conducted using aerobic fitness, demographic, and MetS risk-factor variables in a sample of 2nd- and 3rd-grade children (n = 139) who performed a modified version of a flanker task to assess cognitive control. Flanker performance was also compared between children that met no MetS risk-factor criteria (n = 70), and children who met 1 criterion or more (n = 69). Results: Regression analyses indicated that after controlling for demographic variables and fitness, HDL cholesterol exhibited an independent negative association with flanker reaction time (RT). Group comparisons further revealed that children with no risk factors demonstrated overall shorter RT than the at-risk group. In addition, at-risk children exhibited larger accuracy-interference scores (i.e., poorer performance) for the more difficult conditions of the flanker task that required the up-regulation of cognitive control to meet elevated task demands. Conclusions: These findings are consonant with the previous literature reporting a beneficial influence of aerobic fitness on cognitive control, and reveal new evidence that children without risk factors for MetS exhibit better inhibitory control and increased cognitive flexibility than do at-risk children. In addition to aerobic fitness, these risk factors may serve as important biomarkers for understanding the potential cognitive implications of MetS risk in younger generations.

Keywords: inhibitory control, cognitive flexibility, HDL, aerobic fitness, flanker

Concomitant with the increased prevalence of childhood obesity in the United States has been the increasing risk of metabolic syndrome (MetS) in youth (Duncan, Li, & Zhou, 2004; Saland, 2007). MetS is diagnosed by the presence of three out of five cardiovascular risk factors that may include central obesity, elevated blood pressure (BP), raised triglycerides (TG), reduced high-density lipoprotein cholesterol (HDL), and hyperglycemia (Okosun, Boltri, Lyn, & Davis-Smith, 2010). It is estimated that 6–10% of children and adolescents in the U.S. have MetS, and over 60% have at least one risk factor (de Ferranti, Gauvreau, Ludwig, Newburger, & Rifai, 2006). Emerging evidence suggests that the presence of this adverse metabolic profile not only increases the risk of developing Type 2 diabetes and cardiovascular diseases (CVD), but may also have deleterious consequences for brain health and cognitive function.

Despite the lack of a “gold standard” definition for pediatric MetS (Agatston et al., 2012), the presence of MetS risk factors as early as
5 years of age is predictive of adult Type 2 diabetes (Morrison, Friedman, Wang, & Glueck, 2008). Also during these early years, aspects of children's cognition that are crucial for future school readiness and greater academic performance begin substantial development (Diamond, Barnett, Thomas, & Munro, 2007). As such, the pathological processes associated with adult metabolic dysregulation may have their origins in preadolescent childhood, making early identification of the physical and mental health implications of MetS risk factors a public health priority.

Existing evidence indicates that MetS risk factors may be associated with deficits in aspects of cognition known as cognitive control, which describes a subset of goal-directed, self-regulatory operations encompassing the core processes of inhibition, working memory, and cognitive flexibility (Diamond, 2013). Among older adults, MetS has been associated with an increased risk for cognitive decrements (Taylor & MacQueen, 2007; Yaffe et al., 2004); however, the influence of MetS risk factors on cognitive processes underlying children's scholastic achievement, including cognitive control (Bierman, Nix, Greenberg, Blair, & Domitrovich, 2008), remains unknown. To date, the majority of this health-related research has focused on children's aerobic fitness and weight status, and have revealed that both more fit (Davis & Cooper, 2011; Pontifex et al., 2011) and normal-weight (Kamijo et al., 2012; Li, Dai, Jackson, & Zhang, 2008) children exhibit superior performance on cognitive control tasks than do their peers who are classified as less fit or overweight/obese, respectively. However, only a handful of studies have examined relationships between MetS and its underlying risk-factor components with specific aspects of cognitive control.

Tascilar et al. (2011) recently used neuroimaging (i.e., event-related brain potentials; ERPs) to compare the P3 component using a simple auditory discrimination task in obese and normal-weight children (8–12 years old). ERPs refer to measurable patterns of neuroelectric activation that occur in response to, or in preparation for, stimuli and responses, and the P3, embedded within the ERP, is evoked when individuals attend to or discriminate between stimuli (Polich & Kok, 1995). Tascilar et al. (2011) reported that, in addition to elevated fasting blood-glucose levels (FBG), higher TG, and decreased levels of HDL, obese children exhibited smaller P3 amplitude and longer latency, indicating a decreased ability to allocate attentional resources toward stimulus engagement and delayed stimulus classification and processing speed, respectively (Polich, 2007; Tascilar et al., 2011). We found it intriguing that, when the obese children were divided into groups with or without insulin resistance, further reductions in P3 amplitude and longer latencies were observed in participants with insulin resistance. Additional evidence has since revealed that adolescents (mean age = 17.6 years; range = 14–20) with MetS demonstrate lower performance across a host of cognitive domains, including academic achievement, attention, and cognitive flexibility, compared with control participants with fewer than three risk factors (Yau, Castro, Tagani, Tsui, & Convit, 2012). Brain imaging methods also revealed smaller hippocampal volumes and poorer microstructural integrity of white matter in adolescents with MetS, but perhaps more telling was that the large number of cognitive differences became more pronounced when control participants were limited to only individuals without a single MetS risk factor. As the authors note, decreased cognitive or academic performance and smaller hippocampal volume were observed relative to an increase in the number of MetS risk factors (Yau et al., 2012). Thus, it appears that cognitive differences exist between children with no risk factors and those meeting just one or two criteria for MetS.

These initial studies in children and adolescents with MetS have unveiled cognitive differences that are both consonant with the adult literature and align closely with previous neuroimaging studies reporting differences between higher and lower aerobically fit children (Chaddock et al., 2010a, 2010b; Hillman, Buck, Thamsen, Pontifex, & Castelli, 2009; Hillman, Castelli, & Buck, 2005; Pontifex et al., 2011; Voss et al., 2011). It is important to note, the pattern of results also suggests that children at greater risk for MetS, to varying degrees, exhibit poorer cognitive performance than those without risk factors; however, there has yet to be an investigation of children’s cognitive control that incorporates measures of both aerobic fitness and MetS risk in a single study. Such evidence would be meaningful, bearing in mind that previous studies have shown considerable relatedness between these variables, such that aerobic fitness attenuates MetS risk across multiple Body-Mass-Index (BMI) and weight-class categories in young children (Dubose, Eisenmann, & Donnelly, 2007).

Accordingly, the current study investigated the relationship between MetS risk factors and cognitive control among second- and third-grade children while controlling for aerobic fitness and weight status. We hypothesized that children without risk factors for MetS would demonstrate superior inhibitory control as evidenced by shorter RT and higher accuracy scores during performance of a flanker task. In addition, this pattern of results was expected to be disproportionately greater for the difficult conditions of the task necessitating increased cognitive control, similar to the patterns witnessed in previous investigations of children’s aerobic fitness (Chaddock et al., 2010b; Pontifex et al., 2011; Voss et al., 2011). The identification and explication of the role of MetS risk factors influencing childhood cognitive and brain health is of critical importance for scholastic achievement and cognitive well-being during development and across the life span (Diamond, 2013).

**Method**

**Participants**

Second and third grade children were recruited (by flyers mailed to guardians) from 17 schools in the Lawrence, Kansas area to participate in a larger cluster-randomized trial investigating the effects of physically active academic lessons delivered by classroom teachers on academic outcomes (Donnelly et al., 2013). Recruitment was targeted toward all children, not only those who were overweight or had higher likelihood of meeting the criteria for MetS. The data used for the current study were collected during baseline testing for the larger investigation. Due to a large response, a random sample of students (stratified by grade and sex) from each school was selected from those who provided parental consent/child assent. Guardians completed a demographic questionnaire to provide information about the grade, age, sex, ethnicity, and race of their children, as well as household income. The University of Kansas Medical Center Human Subjects Committee approved the study. Table 1 provides a summary of participant demographics and cardiovascular mean values.
Testing Procedure

Testing for this study was completed over the course of 3 days for each child, with aerobic fitness and cognitive testing taking place on separate occasions. Data collection for the entire sample was completed over the course of approximately 1 month to incorporate all participating schools. All assessments were completed at the children’s respective schools by research staff that were trained and supervised by a qualified coinvestigator. For cognitive testing, each participant was individually removed from their classroom and brought to a quiet area, such as an unused counselor’s office or classroom, which remained free of distractions. Each child was comfortably seated at a large desk where they completed the cognitive control task (described below) using a laptop and a handheld response pad.

Aerobic Fitness Testing

The 20-m Progressive Aerobic Cardiovascular Endurance Run (PACER) subtest of the FitnessGram (Welk & Meredith, 2008) was used to assess children’s aerobic fitness. The PACER has an extensive literature describing the reliability and validity of the test in adults and children (Léger & Lambert, 1982; Léger, Mercier, Gadoury, & Lambert, 1988). Correlations between the PACER and laboratory-based measures of aerobic fitness are moderate to strong, ranging from \( r = .65 \) to \( .83 \) (Mahar, Guerieri, Hanna, & Kemble, 2011; Varness, Carrel, Eickoff, & Allen, 2009). The PACER was administered in small groups, typically 8–12 students, with at least one research staff member per three children, and another staff member overseeing the test. During the PACER, children were instructed to run back and forth between two lines, 20 m apart, as they were paced by a tone on a CD player signaling when he or she should reach the opposite line. The pace began slowly and became progressively faster until the child failed to traverse the 20-m distance in the time allotted on two occasions, at which point the test was over (Carrel et al., 2012; Welk & Meredith, 2008). The greater number of laps a child was able to complete indicated a higher level of aerobic fitness.

MetS Criteria and Group Assignment

Following an overnight fast (12 hr), a blood sample was collected from each participant’s antecubital vein by a licensed phlebotomist. Participants were provided with a breakfast consisting of a bagel, fruit, and juice following the blood draw. Blood samples were immediately transported to the laboratory and processed by centrifuging the samples at \( 70 ^\circ C \) until shipped to the laboratory for analysis. Biomedical measures assessed included FBG, insulin, total cholesterol, low-density lipoprotein cholesterol (LDL), HDL, and TG. FBG, total cholesterol, LDL, HDL, and TG concentrations were measured enzymatically using an Olympus AU400e automated clinical chemistry analyzer (Beckman Coulter, Brea, CA). Insulin levels were measured using a competitive radioimmunoassay manufactured by Linco Research (Millipore, Billerica, MA).

MetS risk-factor criteria were based on cutoffs recommended by de Ferranti et al. (2004), which included HDL (\(< 50 \text{mg/dL}\)), waist circumference (WC; \( > 75\text{th} \) percentile for sex and age; Fryar, Gu, & Ogden, 2012), systolic and/or diastolic BP (\( > 90\text{th} \) percentile for sex, age, and height; Falkner & Daniels, 2004), TG (\( \geq 100 \text{mg/dL}\)), and FBG (\( \geq 110 \text{mg/dL}\)). This pediatric definition was selected because the cutoff value for each risk

<table>
<thead>
<tr>
<th>Measure</th>
<th>No risk</th>
<th>( \geq 1 ) Risk factor</th>
<th>All participants</th>
</tr>
</thead>
<tbody>
<tr>
<td>n</td>
<td>70 (35 female)</td>
<td>69 (43 female)</td>
<td>139 (78 female)</td>
</tr>
<tr>
<td>Age (year)</td>
<td>7.6 (0.5)</td>
<td>7.5 (0.6)</td>
<td>7.6 (0.6)</td>
</tr>
<tr>
<td>Grade</td>
<td>2nd (42.9 %)</td>
<td>(53.6%)</td>
<td>(48.2 %)</td>
</tr>
<tr>
<td></td>
<td>3rd (57.1 %)</td>
<td>(46.4%)</td>
<td>(51.8 %)</td>
</tr>
<tr>
<td>Household income</td>
<td>6.9 (3.5)</td>
<td>6.2 (3.1)</td>
<td>6.6 (3.3)</td>
</tr>
<tr>
<td>BMI (kg/m²)**</td>
<td>16.7 (2.1)</td>
<td>19.5 (4.0)</td>
<td>18.1 (3.4)</td>
</tr>
<tr>
<td>Waist circumference (cm)**</td>
<td>55.4 (4.8)</td>
<td>61.6 (9.9)</td>
<td>58.5 (8.3)</td>
</tr>
<tr>
<td>PACER (no. of laps)**</td>
<td>19.4 (8.5)</td>
<td>15.2 (8.0)</td>
<td>17.3 (8.5)</td>
</tr>
<tr>
<td>Systolic BP (mmHg)**</td>
<td>99.6 (7.0)</td>
<td>104.0 (9.4)</td>
<td>101.8 (8.6)</td>
</tr>
<tr>
<td>Diastolic BP (mmHg)*</td>
<td>57.2 (4.5)</td>
<td>59.2 (5.5)</td>
<td>58.2 (5.1)</td>
</tr>
<tr>
<td>Fasting glucose (mg/dl)</td>
<td>88.5 (6.7)</td>
<td>89.3 (13.6)</td>
<td>88.9 (10.7)</td>
</tr>
<tr>
<td>Triglycerides (mg/dl)**</td>
<td>54.3 (16.4)</td>
<td>73.3 (26.4)</td>
<td>63.7 (23.9)</td>
</tr>
<tr>
<td>Total cholesterol (mg/dl)**</td>
<td>158.3 (19.0)</td>
<td>147.8 (22.3)</td>
<td>153.1 (21.3)</td>
</tr>
<tr>
<td>LDL (mg/dl)</td>
<td>88.4 (18.9)</td>
<td>86.0 (19.2)</td>
<td>87.2 (19.0)</td>
</tr>
<tr>
<td>HDL (mg/dl)**</td>
<td>59.0 (6.3)</td>
<td>47.1 (7.4)</td>
<td>53.1 (9.1)</td>
</tr>
<tr>
<td>BMI percentiles</td>
<td>(No. of children)</td>
<td>(No. of children)</td>
<td>Total (%)</td>
</tr>
<tr>
<td>Underweight (&lt;5th)</td>
<td>1</td>
<td>0</td>
<td>1 (0.7%)</td>
</tr>
<tr>
<td>Normal-weight (5–84th)</td>
<td>55</td>
<td>31</td>
<td>86 (61.9%)</td>
</tr>
<tr>
<td>Overweight (85–94th)</td>
<td>11</td>
<td>14</td>
<td>25 (18.0%)</td>
</tr>
<tr>
<td>Obese (≥95th)</td>
<td>3</td>
<td>24</td>
<td>27 (19.4%)</td>
</tr>
</tbody>
</table>

Note. BMI = Body Mass Index; PACER = Progressive Aerobic Cardiovascular Endurance Run; BP = blood pressure; LDL/HDL = low/high-density lipoprotein cholesterol.

\* \( p < .05 \) between risk factor groups. \** \( p < .01 \).
factor was designed to be closely analogous to the Adult Treatment Panel III (ATPIII; de Ferranti et al., 2004). As de Ferranti and colleagues noted, other proposed pediatric definitions of MetS (of which there is no standard) used restrictive cholesterol and adiposity cutoff points that likely underestimated the prevalence of MetS risk in children. For example, an HDL level of 40 mg/dL represents approximately the 10–25th percentiles in boys and girls versus the 40th percentile cutoff point in adults, as recommended by the ATPIII (de Ferranti et al., 2004). Lastly, children were divided into those with no risk factors ($n = 70$) and those exhibiting one or more of the risk criteria ($n = 69$).

Cognitive-Control Task

Inhibitory control was assessed using a modified Eriksen flanker task (Eriksen & Eriksen, 1979). This task has been previously used to measure children’s ability to inhibit unnecessary or interfering information in a stimulus environment (Mezzacappa, 2004; Pontifex et al., 2011; Voss et al., 2011). The current task required participants to respond to the direction of a centrally presented, child-friendly stimulus, or goldfish “graphic” amid lateral flanking goldfish on each side. Flanking goldfish always faced the same direction; thus, congruent trials resulted whenever the direction of the central and flanking goldfish matched, and incongruent trials occurred when they did not match (please refer to Voss et al., 2011 for stimuli examples). Stimuli were 2-cm tall yellow goldfish, and trials were presented locally for 200 ms on a light blue background with a fixed interstimulus interval (ISI) of 1,700 ms. During the compatible condition, which was completed first, participants were instructed to respond as quickly and accurately as possible according to the direction that the central (i.e., target) goldfish faced. Using their thumbs, participants made a left button press when the central goldfish faced left and a right button press when it faced right. Task difficulty was increased during the incompatible condition by reversing the stimulus-response mappings, such that participants responded opposite the direction of the central, target goldfish. In both conditions, a 40-trial practice block was completed prior to a 100-trial block, both of which were evenly divided among congruent and incongruent trials, with equiprobable directionality of the target stimuli. Participants were excluded from analyses if their mean flanker performance was below chance (i.e., 50%) across compatible and incompatible conditions ($n = 40$), or if a child was reported as having a neurological disorder, ADHD, or a learning disability ($n = 16$). Participants who were absent for school during testing or missing necessary demographic information were also excluded ($n = 11$). Thus, a total of 139 children were incorporated into the final analyses.

Statistical Analysis

The prevalence of MetS was calculated overall and according to sex and weight status (i.e., BMI $\geq$ 85th percentile). Fisher’s exact test was incorporated to identify significant differences in the prevalence of MetS and for each of the individual risk factors across these classifications. Pearson product–moment correlations were used to identify relationships between aerobic fitness and demographic variables, including the number of laps run on the PACER test, BMI, age, grade, sex (coded as $0 = \text{female}$, $1 = \text{male}$), and household income (coded as 1, $<$10,000 per year, through 11, $>100,000$ per year, with $10,000$ increments) using SPSS Version 21 (IBM, Armonk, NY). MetS risk-factor variables were also included, i.e., systolic/diastolic BP, WC, HDL, FBG, and TG. Although LDL is not listed among risk-factor variables for MetS, it was included among the bivariate analyses for possible comparisons with previous investigations (Perry, Stigger, Ainsworth, & Zhang, 2009). In addition, correlations involved the four dependent flanker variables (congruent RT with accuracy, incongruent RT with accuracy), which were included from each response- compatibility manipulation (compatible or incompatible). Finally, flanker accuracy-interference scores were calculated within each compatibility condition (congruent minus incongruent trial accuracy), as well as across compatibilities (compatible minus incompatible accuracy) for each trial type. Thus, larger positive numbers would indicate greater interference (i.e., poorer performance) and thus provide an indication of how children perform as flanker difficulty increases.

Hierarchical regression analyses ($n = 139$) were incorporated to further examine each dependent flanker variable by entering demographic variables into Step 1, aerobic fitness in Step 2, and significant MetS risk factors in Step 3. Grade was included as a demographic factor in these analyses rather than age because it was more strongly associated with all other variables. Assumptions of linearity, equality of variance, independence, and normality were plotted, inspected, and verified using Studentized residuals. Multicollinearity was not observed among any of the independent variables. Accuracy and RT were separately compared between children with ($n = 69$) and without ($n = 70$) risk factors for MetS using a Compatibility (compatible, incompatible) $\times$ Congruency (congruent, incongruent) multivariate analysis of variance (MANOVA) with risk factor (i.e., 0 vs. $\geq 1$) as a between-subjects factor. Flanker accuracy-interference scores were compared using independent-samples $t$ tests. Post hoc comparisons were conducted using the Bonferroni correction, and Cohen’s $d$ is reported to indicate effect size. Significance levels were set at $p = .05$ for all analyses.

Results

Prevalence of MetS Risk

Table 2 displays the prevalence rates of MetS and the individual risk factors according to sex and weight status, the
overall prevalence of MetS was 6.5%, with the most common risk factors being low HDL (40.3%), central obesity (16.5%), and high BP (11.5%). Just over half of the children (50.3%) had two risk factors, 4.3% had three risk factors, and 2.2% had four risk factors (no child had all five risk factors). No differences in MetS risk factors being low HDL (40.3%), central obesity (16.5%), and high BP (11.5%) were observed as a function of sex (ps ≥ .057). As for weight status groups, children who were overweight or obese (≥85th percentile BMI for age) exhibited a higher prevalence of MetS (17.3% vs. 0.0%), elevated TG (13.5% vs. 0.0%), low HDL (59.6% vs. 28.7%), and larger WC (44.2% vs. 0.0%), ps ≤ .001. However, no differences were observed for high BP or FBG (ps ≥ .108).

### Flanker Regression

Initial Pearson product–moment correlations revealed that aerobic fitness was significantly related to WC (r = -.30, p < .001), HDL (r = .31, p < .001), and TG (r = -.28, p = .001). That is, children with higher aerobic fitness had smaller WC, greater levels of HDL, and lower levels of TG. Aerobic fitness was also associated with demographic variables, including BMI (r = -.38, p < .001), grade (r = .19, p = .03), and sex (r = .21, p = .01); a trend for household income (r = .15, p = .08) was observed as well. Thus, more aerobically fit children had lower BMI and were older, male, and living in higher income families. This pattern highlights the need to control for these demographic factors. Regarding the dependent cognitive outcomes, HDL (rs ≥ -.27, ps ≤ .002), TG (rs ≥ .21, ps ≤ .012), and fitness (rs ≥ -.26, ps ≤ .002) all demonstrated significant associations with flanker RT across both trial types in each compatibility condition.

The regression analyses are summarized in Table 3. Correlations suggested no significant relations between LDL or flanker accuracy and any of the other variables; thus, based on the study purpose and the a priori hypotheses, only RT regression analyses are reported below. Demographic variables including grade, sex, household income, and BMI were entered into Step 1 of the analyses. Aerobic fitness was entered at Step 2, based on its significant correlation with flanker RT, and both HDL and TG were entered in Step 3 to determine if the incremental amount of variance accounted for by either risk factor remained significantly associated with flanker RT.

Step 1, which included the demographic variables, did not reach significance for any of the RT measures, ΔR² ≤ 0.059, F(4, 134) ≤ 2.09, ps ≥ .085. In Step 2, fitness accounted for a significant and unique proportion of the variance, ΔR² ≥ 0.038, F(1, 133) ≥ 5.67, ps ≤ .02, across each RT measure, βs ≥ -.23, ps ≤ .022. Step 3 was also significant for compat-

### Table 2

<table>
<thead>
<tr>
<th>Risk factor</th>
<th>Male/female, n (%)</th>
<th>BMI &lt; 85th/BMI ≥ 85th percentile, n (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Metabolic syndromeb</td>
<td>3 (4.9%)/6 (7.7%)</td>
<td>0 (0.0%)/9 (17.3%)</td>
</tr>
<tr>
<td>High blood pressure</td>
<td>8 (13.1%)/8 (10.3%)</td>
<td>7 (8.0%)/9 (17.3%)</td>
</tr>
<tr>
<td>High fasting glucose</td>
<td>0 (0.0%)/2 (2.6%)</td>
<td>1 (1.1%)/1 (1.9%)</td>
</tr>
<tr>
<td>High triglyceridesb</td>
<td>3 (4.9%)/4 (5.1%)</td>
<td>0 (0.0%)/7 (13.5%)</td>
</tr>
<tr>
<td>Low HDLb</td>
<td>19 (31.1%)/37 (47.4%)</td>
<td>25 (28.7%)/31 (59.6%)</td>
</tr>
<tr>
<td>Central obesityb</td>
<td>7 (11.5%)/16 (20.5%)</td>
<td>0 (0.0%)/23 (44.2%)</td>
</tr>
<tr>
<td>Total (n = 139)</td>
<td>61 (44.0%)/78 (56.0%)</td>
<td>87 (63.0%)/52 (37.0%)</td>
</tr>
</tbody>
</table>

Note. BMI = Body Mass Index; HDL = high-density lipoprotein cholesterol.

* p = .057, males vs. females. b p ≤ .001 between weight status groups.

### Table 3

<table>
<thead>
<tr>
<th>Model and variable</th>
<th>flanker compatible</th>
<th>Incongruent</th>
<th>Flanker Incompatible</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>ΔR²</td>
<td>β</td>
<td>ΔR²</td>
</tr>
<tr>
<td><strong>Step 1</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Grade</td>
<td>.035</td>
<td>-.10</td>
<td>-.07</td>
</tr>
<tr>
<td>Sex</td>
<td>-.15</td>
<td>-.13</td>
<td>-.19*</td>
</tr>
<tr>
<td>Household income</td>
<td>-.05</td>
<td>-.04</td>
<td>-.02</td>
</tr>
<tr>
<td>BMI</td>
<td>.02</td>
<td>.04</td>
<td>.11</td>
</tr>
<tr>
<td><strong>Step 2</strong></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>PAGER (no. of laps)</td>
<td>.087**</td>
<td>-.34**</td>
<td>.066**</td>
</tr>
<tr>
<td><strong>Step 3</strong></td>
<td>.068**</td>
<td>-.21*</td>
<td>.058*</td>
</tr>
<tr>
<td>HDL</td>
<td>.15</td>
<td>.11</td>
<td>.09</td>
</tr>
</tbody>
</table>

Note. RT = response time; BMI = Body Mass Index; PAGER = Progressive Aerobic Cardiovascular Endurance Run; HDL = high-density lipoprotein.

* p = .055.

' p < .05. ** p < .01.
ible flanker RTs after controlling for fitness, $\Delta R^2 \geq 0.058$, $F(2, 131) \geq 4.50$, $p \leq .013$, yet it appeared that this effect was primarily driven by the inclusion of HDL, $\beta_s \approx -.21$, $p \leq .023$, as opposed to TG, $\beta_s \leq .15$, $p \geq .106$. As for incompatible RT, similar effects were observed with Step 3 explaining a unique proportion of variance for incongruent trials, $\Delta R^2 = 0.044$, $F(2, 131) = 3.36$, $p = .038$, and an effect for HDL approached significance, $\beta = -.18$, $p = .055$. Alternatively, Step 3 did not reach significance for congruent trials, $\Delta R^2 = 0.035$, $F(2, 131) = 2.68$, $p = .072$. As such, both higher fit children and individuals with higher HDL had shorter RT for incongruent trial performance in particular. The partial regression plots are displayed in Figure 1.

**Group Comparison**

In addition to the demographic variables controlled for in Step 1 of the regression analyses, aerobic fitness was also entered as a covariate into each MANOVA when comparing flanker performance between children with and without risk factors for MetS.

**RT.** A main effect for risk factor, $F(1, 132) = 5.87$, $p = .017$, $\eta^2 = .04$, was observed, indicating that children with no risk factors ($M = 514.5$, $SE = 13.6$ ms) displayed shorter RT than those with one or more risk factors ($M = 562.5$, $SE = 13.7$ ms; see Figure 2A), $d = 0.5$. No other effects involving compatibility or congruency were observed, $F(1, 132) \leq 0.90$, $p \geq .345$, $\eta^2 \leq .01$.

**Accuracy.** For flanker accuracy, an interaction of Risk Factor $\times$ Compatibility $\times$ Congruency, $F(1, 132) = 10.73$, $p = .001$, $\eta^2 = .08$, was observed. Given that accuracy did not differ by group on any of the flanker trial types, $t(137) = 1.20$, $p = .232$, $d = 0.2$, separate 2 (compatibility) $\times$ 2 (congruency) MANOVAs were conducted within each group. A main effect of congruency, $F(1, 132) = 38.62$, $p < .001$, $\eta^2 = .36$, was uncovered in both groups, which was superseded by a Compatibility $\times$ Congruency interaction, $F(1, 169) = 16.99$, $p < .001$, $\eta^2 = .20$. First, analyzing compatible accuracy revealed that both groups exhibited poorer performance for incongruent trials (no risk: $M = 65.1\%$, $SE = 1.8$; one risk factor: $M = 67.9\%$, $SE = 1.6$) compared with congruent trials (no

![Figure 1.](image-url) Partial regression plots ($n = 139$) depicting the relation between fitness and A) compatible, and B) incompatible flanker RT. Plots (C) and (D) display this same association between HDL levels and flanker RT. Partial correlation ($pr$) values are provided. **$^*$** $p < .01$, **$^*$** $p < .05$, **$^*$** $p = .055$. 

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risk: $M = 75.5\%$, $SE = 1.7$; ≥ one risk factor: $M = 76.4\%$, $SE = 1.6$), $t(8.74, p < .001, d = 0.7$. However, this same pattern (incongruent < congruent accuracy) was only observed for the incompatible manipulation in the group with ≥ one risk factor (incongruent: $M = 69.9\%$, $SE = 2.0$; congruent: $M = 73.3\%$, $SE = 1.8$), $t(68) = 3.27, p = .002, d = 0.2$, compared with children with no risk factors (incongruent: $M = 72.6\%$, $SE = 1.9$; congruent: $M = 72.8\%$, $SE = 2.0$), who maintained equal performance on both trial types, $t(69) = 0.17, p = .862, d < 0.1$. As for analyzing each trial type (i.e., congruent, incongruent) across compatibility condition, children with ≥ one risk factor demonstrated decreased congruent trial accuracy as difficulty increased during the incompatible manipulation, $t(68) = 2.24, p = .025, d = 0.2$. Their incongruent trial accuracy increased slightly, yet this difference was not significant, $t(68) = 1.36, p = .178, d = 0.1$. On the other hand, children without risk factors exhibited no significant increase in congruent trial accuracy, $t(69) = 1.40, p = .165, d = 0.2$, and the same trend of increased incongruent trial accuracy was significant in this group of children, $t(69) = 4.28, p < .001, d = 0.5$.

**Accuracy interference.** To more accurately portray these group differences, independent samples $t$ tests were conducted on interference scores that were calculated within each compatibility condition (congruent–incongruent accuracy), as well as across compatibilities for each trial type (compatible–incompatible accuracy). As depicted in Figure 2B, both groups exhibited large congruency-interference scores during the compatible condition, $t(137) = 1.40, p = .178, d = 0.2$. In the incompatible condition, children with ≥ one risk factor exhibited significantly greater congruency interference than children with no risk, $t(137) = 2.22, p = .025, d = 0.4$. Lastly, groups had comparable interference scores for congruent trials across compatibilities, $t(137) = 0.15, p = .882, d < 0.1$; however, at-risk children demonstrated greater interference for incongruent trials, $t(137) = 2.42, p = .017, d = 0.4$, than did children with no risk factors.

![Figure 2](image1.png)

**Discussion**

These data have revealed that children meeting no risk criteria for MetS exhibit superior performance on a widely implemented task of inhibitory control, which remained even after controlling for aerobic fitness level. As hypothesized, children without risk factors for MetS demonstrated greater inhibitory control and cognitive flexibility, as denoted by overall shorter flanker RTs and less interference during the most difficult portions of the task. This pattern of results closely matches previous investigations of children’s aerobic fitness (Chaddock et al., 2010b; Pontifex et al., 2011; Voss et al., 2011), which also had a significant and independent influence on flanker RT in the current study. These findings extend prior literature by revealing that individual aspects of children’s cardiovascular health have distinct associations with aspects of cognitive control, which may have considerable implications for children’s health promotion, learning, and scholastic achievement.

Although the evidence is limited, the findings are consistent with previous studies that have highlighted the importance of assessing MetS risk factors when studying cognitive function and brain health in children. Among the biomedical measures, HDL and TG were initially correlated with flanker RT; however, regression analyses demonstrated that only HDL was independently associated with RT. These findings are in contrast to those observed by Perry et al. (2009), who uncovered no links between children’s (6–16 years old) cholesterol levels and cognitive or academic performance using data from the National Health and Nutrition Examination Survey. However, to our knowledge, the current study and Perry et al. (2009) are the only two studies that have assessed the relationship between cholesterol and cognitive function in preadolescent children. In addition, the studies differ in their assessments of cognitive function, as the present study focused on inhibitory control and cognitive flexibility, whereas Perry et al. (2009) examined academic achievement and aspects of memory. Another expla-
nation for the discordance observed may have been the fasting compliance prior to blood-sample collection in Perry et al. (2009), which was only required of the children over the age of 12 years. Thus, one of the strengths of the study presented herein is the reliance on fasting blood measures of MetS. However, despite the significant association between HDL and RT in the current study (as well as the overall group differences), it is unclear how meaningful the RT findings are, given that Perry et al. (2009) failed to find any association between HDL and cognitive or academic performance. Future research will need to pursue this question to determine whether this pattern of results is meaningful. All told, the influence of children’s cholesterol levels on cognitive performance appears to be one area of research with considerable promise that warrants future inquiry.

Another important finding of this study was the observed differences in flanker performance as a function of MetS risk-factor presence. Children with no risk factors exhibited overall shorter RTs, and less accuracy interference when flanker task demands were elevated. These patterns of greater cognitive flexibility when faced with greater cognitive demands are comparable to those previously witnessed in studies of children’s aerobic fitness. Neuroimaging findings from these studies have revealed that behavioral performance differences observed between higher and lower fit children may relate to differential patterns of brain activation (Pontifex et al., 2011; Voss et al., 2011), as well as the underlying integrity of specific brain regions, including the basal ganglia (Chaddock et al., 2010b). However, it remains unknown whether aerobic fitness and HDL share similar mechanisms that influence cognitive control in childhood. Nevertheless, the current findings point to the possible future utility of both aerobic fitness and HDL levels in predicting cognitive control strategies in prepubertal children.

One particular aspect of research deserving much attention is determining the mechanisms that might be responsible for eliciting these cognitive differences in children with MetS. Cholesterol plays a critical role in the central nervous system by possibly activating signaling pathways that trigger synaptogenesis, and may be necessary for supporting the structural demands of this process (Barres & Smith, 2001; Göritz, Mauch, Nägler, & Pfrieger, 2002). Although the evidence is limited among children, studies in older adults have linked low HDL to smaller hippocampal volume (Wolf et al., 2004), low gray-matter volume (Ward et al., 2010), and cognitive impairment (van Exel et al., 2002). Alternatively, the relationship between HDL and cognitive function observed in this study may be a result of unmeasured variability in inflammatory markers such as C-reactive protein (CRP) and interleukin-6, which have been reported a prospective design to show that older individuals with MetS exhibited greater cognitive decline over a 4-year period, but only for participants with high levels of these inflammatory markers. Therefore, future investigations of MetS and cognition in children should incorporate measures to assess these inflammatory markers and attempt to determine their influence.

There are undoubtedly other explanations that may be considered when discussing the present results. There have been several attempts to isolate the potential genetic underpinnings of MetS, which are believed to exist from evidence demonstrating that insulin resistance and other risk factors tend to cluster in families (Group, 2000). Although there has been some discussion of a common underlying genetic factor rooted around central adiposity, much of the data would suggest that it is an interaction between various genes that accounts for what is collectively referred to as MetS (Joy, Lahiry, Pollex, & Hegele, 2008). As such, there may be reason to suspect that variations in these genes could also influence aspects of cognition or level of aerobic fitness. In any case, many researchers have acknowledged the key role that environmental factors play in developing MetS. As Roche et al. (2005) noted, the human genome has changed very little within the last 10 years, yet the prevalence of MetS and specific risk factors have increased exponentially. Accordingly, it is highly plausible that modifying any one of these variables, especially during a time when humans are undergoing development, may result in widespread changes that are observable throughout the body and brain.

The current study has several limitations. Given the cross-sectional design, these findings do not warrant causal inference of MetS risk or aerobic fitness on children’s cognitive control. Therefore, prospective studies are needed to monitor and assess longitudinal changes in MetS risk factors and cognitive performance. The adopted definition of MetS (de Ferranti et al., 2004) also partially restricts comparisons with other studies. However, this has been a common limitation throughout the pediatric literature, as there is no consensus definition of MetS in childhood (Zimmet et al., 2007). Many definitions have, however, revealed that the risk of MetS varies with regard to a child’s race or ethnicity (e.g., de Ferranti et al., 2004), which indicates that it would be advantageous for future studies to include large, racially diverse samples. Thus, the group of predominately White children used in this investigation may limit the generalizability of the current findings to others. The aforementioned limitations notwithstanding, the major strengths of this study included evaluation of MetS risk using biomedical measures, aerobic fitness using a reliable field estimate, and utilization of a popular cognitive task known to tap specific aspects of cognitive control.

Knowing that the presence of MetS risk factors during childhood predicts the onset of adult MetS (Morrison et al., 2008), we stress, with the current findings, the importance of preventing this risk in young children to promote better brain health and cognitive function starting early in life. The present results also add to the growing literature by suggesting that lower levels of HDL in childhood may detrimentally influence cognitive control. Considering that both HDL and aerobic fitness (variables shown here to be independently associated with inhibitory control) are susceptible to modulation by regular physical activity (Kamijo et al., 2011; Strong et al., 2005), the present

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3 White children accounted for 79.1% of participants; the remaining distribution was divided among African American (3.6%), Asian (0.7%), American Indian/Alaska Native (1.4%), two or more races (13.7%), and those who withheld their information (n = 2).
findings are consistent with previous work highlighting the need to maintain an active lifestyle to achieve optimal cognitive function, beginning in childhood (Institute of Medicine of the National Academies, 2013). For this area of research to continue progressing, one of the most important accomplishments will be establishing a common pediatric definition of MetS, similar to the ATP III for adults. In addition to a common definition, future research will benefit from the inclusion of measures of inflammatory markers while continuing to monitor levels of aerobic fitness and weight status. By delineating which of these health-related variables are most influential for children’s cognitive health, school administrators and other professionals will gain considerable insight for developing future comprehensive health recommendations intended to foster improved academic and cognitive performance and overall well-being in children.

References


